# **Coupled Hidden Markov Models: some computational challenges**

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### 1. CHMMs

Coupled Hidden Markov Models (CHMMs) are a natural extension of HMMs when there are multiple observation sequences with dependencies:



### 2. The naïve approach

One might naïvely reformulate as:



Objective is inference in the minimal setting of  $N = 10, C = 100, T = 10^5$ . This leads to numerous challenges:

- computing forward variable  $\implies TN^C$ additions and TC multiplications;  $\ge 10^{105}$  elementary operations
- forward variable requires 8TN<sup>C</sup> bytes of memory to store; ≥ 7.45 × 10<sup>96</sup> GB memory
   transition matrix is N<sup>C</sup> × N<sup>C</sup>. ≥ 9.31 × 10<sup>190</sup> GB memory

**Notation:**  $x_t^{(i)}, y_t^{(i)}$  : hidden state/obs at time *t* in chain *i*.

 $Z_{t} = (X_{t}^{(1)}, \dots, X_{t}^{(3)}) \text{ and } Y_{t}^{(i)} | Z_{t} = Y_{t}^{(i)} | X_{t}^{(i)}$   $\implies \text{ for } C \text{ chains with } X_{t}^{(i)} \in \{1, \dots, N\},$  $|Z_{t}| = N^{C}.$ 

Natural forward variable becomes:

Hence, naïve approach clearly a non-starter.

$$\alpha_t(i_1, \dots, i_C) = \mathbb{P}(Y_{1:t}^{(1:C)} = y_{1:t}^{(1:C)}, X_t^{(1:C)} = i_{1:C})$$

$$= \left(\sum_{j_1=1}^N \dots \sum_{j_C=1}^N \alpha_{t-1}(j_1, \dots, j_C) \prod_{k=1}^C \mathbb{P}(X_t^{(k)} = i_k \mid x_{t-1}^{(1:C)} = j_{1:C})\right) \prod_{k=1}^C f_{Y \mid X}(y_t^{(k)} \mid i_k)$$

# 3. Existing approaches

Saul and Jordan (1999) Mixture Model

$$\mathbb{P}(X_t^{(i)} \mid x_{t-1}^{(1:C)}) = \sum_{k=1}^{C} \omega_{ki} \mathbb{P}(X_t^{(i)} \mid x_{t-1}^{(k)})$$

 $\omega_{ki}$  can be viewed as mixing weights, or strength of effect of chain *k* on chain *i*. Now only  $NC^2$  parameters.

#### Zhong and Ghosh (2002) Marginal Composite

## 5. Our main interest and scaling towards C = 100 — initial work

There is some interest in inference on model parameters, but our *primary* interest is actually in inferring dependence structure. e.g. in genomics data set this could infer ancestry. ∴ direct multinomial logistic regression transition model: a blocked spike-and-slab prior for Bayesian variable selection is then equivalent to inferring the hidden layer structure.

Then sample  $\mathbf{X}_{1:T}^{(i)} | \boldsymbol{\beta}, \boldsymbol{\lambda}, \mathbf{Y}_{1:T}^{(i)}, \mathbf{X}_{1:T}^{(-i)}$  backwards, since:

$$\mathbb{P}(X_T^{(l)} = j \mid \mathbf{y}_{1:T}^{(l)}, \mathbf{x}_{1:T}^{(-l)}) = \sum_{i=1}^{l} \alpha_{Tij}^{(l)}$$

 $\mathbb{P}(X_{T-t}^{(l)} = i \mid X_{T-t+1}^{(l)} = j, \mathbf{y}_{1:T}^{(l)}, \mathbf{x}_{1:T}^{(-l)}) \propto \alpha_{(n-t+1)ij}^{(l)}$ 

#### Logistic regression

Currently using Holmes and Held (2006).

Likelihood  $\mathbb{P}(Y_{1:T}^{(1:C)}) \approx \prod_{k=1}^{C} \mathbb{P}(Y_{1:T}^{(k)}) = \prod_{k=1}^{C} \sum_{i=1}^{N} \alpha_T^{(k)}(i)$ 

with  $\alpha_T^{(k)}(i)$  itself a factored approximation of the forward variable. Only C = 2 example.

**Sherlock** *et al.* **(2013) Structured Transitions** Uses structured transition matrix for each chain, where probabilities modelled with a logistic regression with others chains (and external factors) as covariates.

**Choi** *et al.* (2013) Logistic Regression Similarly, a transition matrix per chain, with logistic regression transition probabilities. But, for speed, ad-hoc inferential procedure: mixture model EM to infer observation model, Viterbi to select most likely hidden sequence, IRLS on subsample to fit LR with lasso+AIC.  $C = 39, N = 2, T = 15.4 \times 10^6$ 

#### MCMC sampler

- Hidden states: conditional forward/ stochastic-backward X<sup>(i)</sup><sub>1:T</sub> | β, λ, Y<sup>(i)</sup><sub>1:T</sub>, X<sup>(-i)</sup><sub>1:T</sub> for i ∈ {1,...,C}
   Multinomial logistic parameters
  - $oldsymbol{eta} \mid \mathbf{X}_{1:T}^{(1:C)}$
- Observation model parameters  $\boldsymbol{\lambda} \mid \mathbf{Y}_{1:T}^{(1:C)}, \mathbf{X}_{1:T}^{(1:C)}$

**Hidden states** Define conditional forward variable

$$\begin{split} {}^{(l)}_{tjk} &= \mathbb{P}(y_t^{(l)}, X_{t-1}^{(l)} = j, X_t^{(l)} = k \,|\, \mathbf{y}_{1:t-1}^{(l)}, \mathbf{x}_{1:T}^{(-l)}) \\ &= \left(\sum_{i=1}^N \alpha_{(t-1)ij}^{(l)}\right) \frac{\exp(\tilde{\mathbf{x}}_{t-1}^{*j} \boldsymbol{\beta}_k^{(l)})}{1 + \sum_{n=1}^{N-1} \exp(\tilde{\mathbf{x}}_{t-1}^{*j} \boldsymbol{\beta}_n^{(l)})} \\ &\times f_{Y_t^{(i)} \,|\, X_t^{(i)}} \left(y_t^{(i)} \,|\, k\right) \end{split}$$

Results



#### 4. Our hidden layer model

#### 6. Current work



#### **Probit regression**

 $\alpha$ 

Adapting Pakman and Paninski (2013), a Hamiltonian Monte Carlo sampler for truncated multivariate Gaussian and binary distributions. Achieved substantial speedup vs author's reference C++ implementation by exploiting problem specific features.

Currently exploring GPU implementation:

boundary hit times embarassingly parallel; minimum hit time a reduction operator; entire problem can propagate on GPU. **Hidden states** 

Also, exploring block sampling hidden states. Need to find an algorithm to partition chains in some sense 'optimally': mixing -vs- compute.

### References

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www.i-like.org.uk